

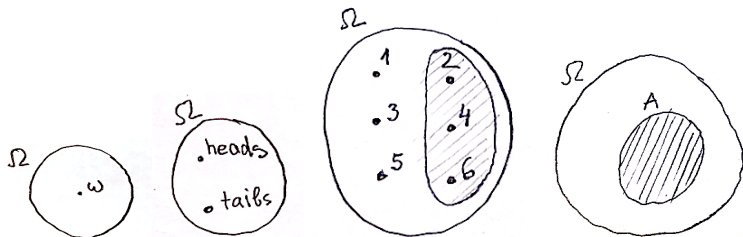
Theory of Unbreakable Ciphers

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Sample Space and Events

Ω -sample space, that contains all possible outcomes $\omega \in \Omega$.



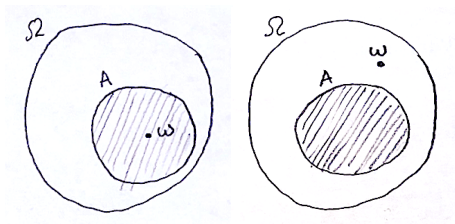
For example, $\Omega = \{\text{heads}, \text{tails}\}$ for a coin, and $\Omega = \{1, \dots, 6\}$ for a die.

Events are subsets $A \subseteq \Omega$.

For a die, the event $\{2, 4, 6\}$ means that the outcome is even.

When do Events Happen?

An event A *happens* if $\omega \in A$ for the actual outcome ω .



Empty event \emptyset is called the *impossible event* (it *never* happens)

Ω is called the *universal event* (it *always* happens)

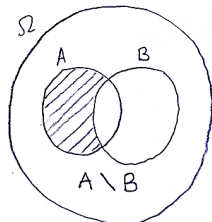
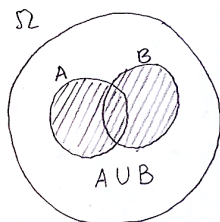
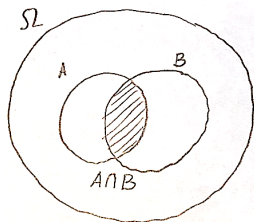
Operations with Events

For every two events A and B we can compute:

Intersection A and B $A \cap B = \{\omega \in \Omega: \omega \in A \text{ and } \omega \in B\}$

Union A or B $A \cup B = \{\omega \in \Omega: \omega \in A \text{ or } \omega \in B\}$

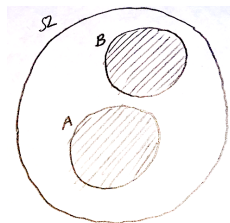
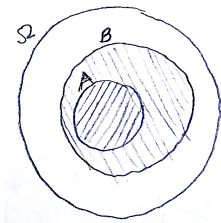
Difference A but not B $A \setminus B = \{\omega \in \Omega: \omega \in A \text{ and } \omega \notin B\}$



Relations Between Events

Inclusion: Event A *implies* event B , if $A \subseteq B$, i.e. if $\omega \in A$ always implies $\omega \in B$. If A happens then B happens.

Exclusion: Events A and B are *mutually exclusive* if $A \cap B = \emptyset$, i.e. A and B cannot simultaneously happen.



Some Properties

Theorem (1)

$$A = (A \setminus B) \cup (A \cap B)$$

Proof.

We prove (a) $A \subseteq (A \setminus B) \cup (A \cap B)$ and (b) $(A \setminus B) \cup (A \cap B) \subseteq A$

(a) If $\omega \in A$ then either:

- $\omega \in B$, which implies $\omega \in A \cap B$, or
- $\omega \notin B$, which implies $\omega \in A \setminus B$

(b) If $\omega \in (A \setminus B) \cup (A \cap B)$, then either:

- $\omega \in A \setminus B$, which implies $\omega \in A$, or
- $\omega \in A \cap B$, which also implies $\omega \in A$



Some Properties

Theorem (2)

$$A \cup B = (A \setminus B) \cup B$$

Proof.

We prove (a) $A \cup B \subseteq (A \setminus B) \cup B$ and (b) $(A \setminus B) \cup B \subseteq A \cup B$

(a) If $\omega \in A \cup B$, then either:

- $\omega \in B$ or
- $\omega \notin B$ and $\omega \in A$, which implies $\omega \in A \setminus B$.

(b) If $\omega \in (A \setminus B) \cup B$ then either:

- $\omega \in B$ or
- $\omega \in A \setminus B$ that implies $\omega \in A$.



Event Algebra

The set \mathcal{F} of all events we consider must be a *sigma-algebra*:

- $\Omega \in \mathcal{F}$
- If $A \in \mathcal{F}$, then $\Omega \setminus A \in \mathcal{F}$
- If $A_1, A_2, A_3, \dots \in \mathcal{F}$, then $A_1 \cup A_2 \cup A_3 \cup \dots \in \mathcal{F}$

If $A \in \mathcal{F}$, then A is said to be a *measurable* subset.

Example: The *set* $P(\Omega)$ *of all subsets* of Ω is a sigma-algebra.

In this class, we mostly assume that $\mathcal{F} = P(\Omega)$.

Probability Measure

Probability (measure) is a function $P: \mathcal{F} \rightarrow \mathbb{R}$ such that:

- *PM1*: $0 \leq P[A] \leq 1$ for any event $A \in \mathcal{F}$.
- *PM2*: $P[\Omega] = 1$
- *PM3*: If $A_1, A_2, \dots \in \mathcal{F}$ are mutually exclusive, then

$$P[A_1 \cup A_2 \cup \dots] = P[A_1] + P[A_2] + \dots$$

The triple (Ω, \mathcal{F}, P) is called a *probability space*.

If \mathcal{F} is the set of all subsets of Ω , we omit \mathcal{F} and say that a probability space is a pair (Ω, P) .

Some Implications

Theorem

$$P[\Omega \setminus A] = 1 - P[A]$$

Proof.

By *PM2*, we have $P[\Omega] = 1$. As A and $\Omega \setminus A$ are mutually exclusive, and $(\Omega \setminus A) \cup A = \Omega$, by *PM3*, we have $P[\Omega \setminus A] + P[A] = P[\Omega] = 1$ and hence

$$P[\Omega \setminus A] = \underbrace{P[\Omega \setminus A] + P[A]}_1 - P[A] = 1 - P[A] .$$



Some Implications

Theorem

$$P[A] + P[B] = P[A \cap B] + P[A \cup B]$$

Proof.

By Thm. 1: $A = (A \setminus B) \cup (A \cap B)$. As $A \setminus B$ and $A \cap B$ are mutually exclusive, by **PM3**: $P[A] = P[A \setminus B] + P[A \cap B]$. Hence,

$$P[A] + P[B] = P[A \setminus B] + P[B] + P[A \cap B]$$

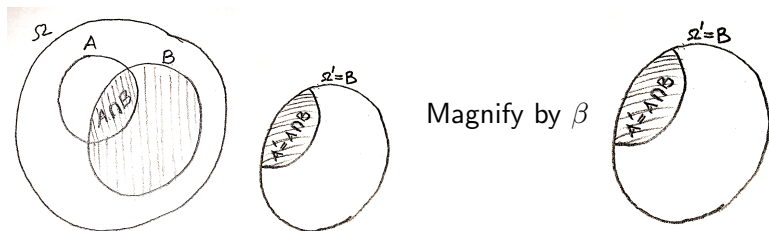
By Thm. 2: $A \cup B = (A \setminus B) \cup B$. As $A \setminus B$ and B are mutually exclusive, by **PM3**: $P[A \cup B] = P[A \setminus B] + P[B]$. Hence,

$$P[A] + P[B] = \underbrace{P[A \setminus B] + P[B]}_{P[A \cup B]} + P[A \cap B] = P[A \cup B] + P[A \cap B] .$$



Learning

Somehow we learn that an event B (with $P[B] \neq 0$) happens, i.e. $\omega \in B$. Probability space (Ω, P) collapses to a new space (Ω', P') , where $\Omega' = B$.



We want that there is β , so that $P'[A] = \beta \cdot P[A \cap B]$ for any event A .
 As in the new space, $P'[B] = P'[\Omega'] = 1$, we have $\beta = \frac{1}{P[B \cap B]} = \frac{1}{P[B]}$, i.e.

$$P'[A] = \frac{P[A \cap B]}{P[B]} .$$

Conditional Probability

The probability

$$P'[A] = \frac{P[A \cap B]}{P[B]}$$

is denoted by $P[A | B]$ and is called the *conditional probability* of A assuming that B happens, i.e.

$$P[A | B] = \frac{P[A \cap B]}{P[B]}$$

Corollary (Chain Rule):

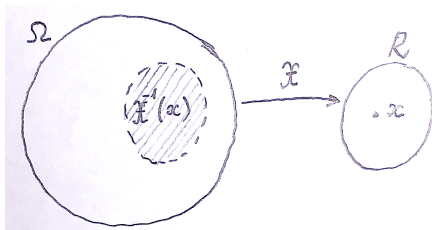
$$P[A \cap B] = P[B] \cdot P[A | B] = P[A] \cdot P[B | A]$$

Random Variables

Random variable X is any function $X: \Omega \rightarrow R$, where R is called the *range* of X . We write R_X to denote the range of X

For any $x \in R$, we define $X^{-1}(x)$ as the event $\{\omega: X(\omega) = x\}$ and use the notation:

$$P_X[x] = P[X = x] = P[X^{-1}(x)] .$$



Finite Range Random Variables

In cryptography, we mostly assume that the range R is *finite*.

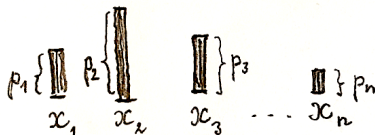
Note that if $x \neq x'$, then the events $X^{-1}(x)$ and $X^{-1}(x')$ are mutually exclusive and as $\cup_{x \in R} X^{-1}(x) = \Omega$, we have:

$$\sum_x \mathbb{P}_X[x] = \mathbb{P}[\cup_{x \in R} X^{-1}(x)] = \mathbb{P}[\Omega] = 1 .$$

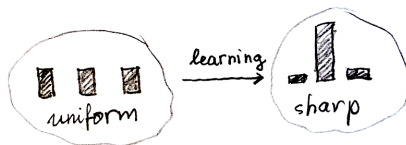
Probability Distributions and Histograms

Assume R is finite and $R = \{x_1, x_2, \dots, x_n\}$.

The sequence of values (p_1, p_2, \dots, p_n) , where $p_i = P_X[x_i]$, is called the *probability distribution* of X .



Histograms are graphical representations of probability distributions.



Independent Events and Random Variables

Events A and B are said to be *independent* if $P[A \cap B] = P[A] \cdot P[B]$

If $P[A] \neq 0 \neq P[B]$, then independence is equivalent to:

$$P[A | B] = P[A] \quad \text{and} \quad P[B | A] = P[B] ,$$

i.e. the probability of A does not change, if we learn that B happened.

We say that X and Y are *independent random variables* if for every $x \in R_X$ and $y \in R_Y$:

$$\begin{aligned} P[X = x, Y = y] &= P[X^{-1}(x) \cap Y^{-1}(y)] = P[X^{-1}(x)] \cdot P[Y^{-1}(y)] \\ &= P[X = x] \cdot P[Y = y] . \end{aligned}$$

This means that the probability distribution of X does not change, if we learn the value of Y , and vice versa

Direct Product of Random Variables

By the *direct product* XY (or (X, Y)) of random variables X and Y on a probability space (Ω, \mathcal{P}) is a random variable defined by

$$(XY)(\omega) = (X(\omega), Y(\omega)) .$$

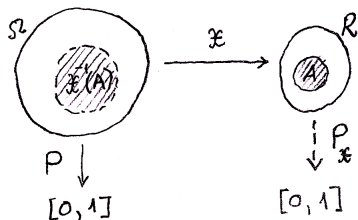
Factor Space

Let X be a random variable (with range R) on a probability space (Ω, P) .

If we take $\Omega' = R$ and define a probability function P_X on R as follows:

$$P_X[A] = P[X^{-1}(A)]$$

where $X^{-1}(A) = \{\omega \in \Omega : X(\omega) \in A\}$, we get a probability space (R, P_X) that we call a *factor space*.



Probabilistic Model of a Cipher

Plaintext X , key Z and ciphertext $Y = E_Z(X)$ are random variables on (Ω, \mathcal{P}) . It is mostly assumed that X and Z are independent.

As we need only X , Y , and Z , we study the *factor space* (R_{XZ}, P_{XZ}) that consists of all possible plaintext-key pairs (x, z) , whereas

$$\mathbf{P}_{XZ}[x, z] = \mathbf{P}[X = x] \cdot \mathbf{P}[Z = z] = p(x) \cdot p(z)$$

$X(x, z) = x$, $Z(x, z) = z$, and $Y(x, z) = E_z(x)$.

Some Observations

$$\begin{aligned}
 p(y) &= \mathbf{P}_{XZ}[Y = y] = \sum_{x,z} \mathbf{P}[x, z][E_z(x) = y] \\
 &= \sum_x p(x) \sum_z p(z)[E_z(x) = y] \\
 p(x, y) &= \mathbf{P}_{XZ}[X = x, Y = y] = \sum_z \mathbf{P}[x, z][E_z(x) = y] \\
 &= p(x) \sum_z p(z)[E_z(x) = y]
 \end{aligned}$$

Here, $[A(x, yz)]$ is the so-called *Iverson symbol*:

$$[A(x, y, z)] = \begin{cases} 1 & \text{if } A(x, y, z) \text{ holds} \\ 0 & \text{otherwise} \end{cases}$$

Definition of Unbreakable Cipher

A cipher is *unbreakable* if ciphertext Y and plaintext X are independent.

Theorem

If Z is independent of X , Z is uniformly distributed and for every plaintext x and for every ciphertext y there is a unique key z such that $E_z(x) = y$, then the cipher is unbreakable.

Proof.

Due to the unique z , we have $\sum_z p(z)[E_z(x) = y] = p(z)$, and thus

$$\begin{aligned} p(x | y) &= \frac{p(x, y)}{p(y)} = \frac{p(x) \sum_z p(z)[E_z(x) = y]}{\sum_x p(x) \sum_z p(z)[E_z(x) = y]} = \frac{p(x)p(z)}{p(z) \sum_x p(x)} \\ &= \frac{p(x)p(z)}{p(z) \cdot 1} = p(x) \end{aligned}$$



Shift Cipher in Unbreakable

Shift cipher: $y = E_z(x) = x + z \pmod{m}$

For every x and y , there is one and only one z , such that $E_z(x) = y$:

$$z = y - x \pmod{m} .$$

Therefore, by the theorem above, shift cipher is unbreakable.

Redundancy of English

In case of 26-letter alphabet, a single letter contains $\log_2 26 \approx 4.7$ bits of information.

Random n -letter sequence contains $4.7n$ bits of information.

Meaningful english texts contain just about 1.5 bits of information per letter.

There are $2^{4.7n}$ arbitrary n -letter sequences, $2^{1.5n}$ of them meaningful

The probability that a randomly chosen n -letter sequence is meaningful is:

$$\mu = \frac{2^{1.5n}}{2^{4.7n}} = 2^{-3.2n} .$$

Exhaustive Key Search

Given a ciphertext y

For all keys z , check if $D_z(y)$ is a meaningful text

Success, if there is just one z for which $D_z(y)$ is meaningful

Ideal Cipher Model

For every key z , the function $E_z: \mathbf{X} \rightarrow \mathbf{Y}$ is a randomly chosen one-to-one function

This implies that the decryption function $D_z: \mathbf{Y} \rightarrow \mathbf{X}$ is also a randomly chosen one-to-one function

If $z_1 \neq z_2$, then $X_1 = D_{z_1}(y)$ and $X_2 = D_{z_2}(y)$ are independent uniformly distributed random variables

Unicity Distance

Unicity distance: message length n_0 for which the plaintext can be derived from the ciphertext via exhaustive key search

Let y be a ciphertext

Assume there are 2^k possible keys z , one of which is the right key

The probability that $D_z(y)$ is meaningful for a fixed wrong key z is $\mu = 2^{-3.2n}$

The probability that $D_z(y)$ is meaningful for any of the wrong keys is bounded by $(2^k - 1)\mu$ and also by $2^k\mu = 2^{k-3.2n}$

If $n > n_0 = \frac{k}{3.2}$, the success probability of exhaustive search increases rapidly

Unicity Distance for Substitution Ciphers

The number of keys is $26!$

Hence, $k = \log_2(26!) \approx 88.4$

Therefore, the unicity distance is $n_0 = 88.4/3.2 \approx 28$